

# **STAT537: Statistics for Research I: Midterm**

Due on November 3, 2016

*Dr. Schmidhammer TR 11:10am – 12:25pm*

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## Problem 1

Potencies dataset

*Solution.* (a) **Create a stem-and-leaf plot for these data.**

The decimal point is at the |

```

22 | 79
23 | 0234
23 | 68
24 | 013
24 | 589
25 | 0244
25 | 89
26 | 144
26 | 7799
27 | 123

```

- (b) **Assess the normality of these data.** The following Shapiro-Wilk normality indicates that the p-value is  $0.08275 > 0.05$ . Hence we do not have enough information to reject the  $H_0$ . Therefore, we may consider this data obeys normal distribution. And the corresponding Normal QQ plot can be found in Figure.1 which confirms our conclusion.

```
> shapiro.test(potencies);
```

Shapiro-Wilk normality test

data: potencies

W = 0.93847, p-value = 0.08275

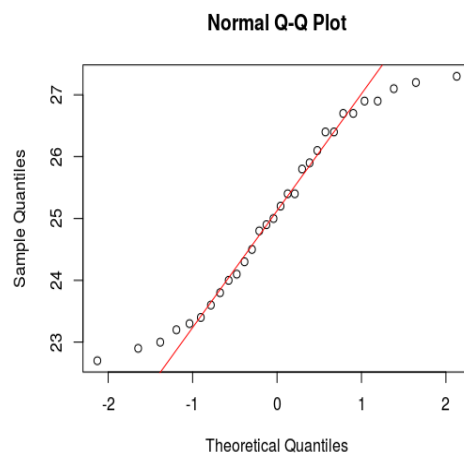


Figure 1: Normal QQ plot.

- (c) **Provide a 99% Confidence Interval for the average potency.** Since we do not know the variance of the population, hence the formula of the 99% ( $\alpha = 0.01$ ) Confidence Interval is

$$\left[ \mu - t_{\alpha/2} \left( \frac{\sigma}{\sqrt{n}} \right), \mu + t_{\alpha/2} \left( \frac{\sigma}{\sqrt{n}} \right) \right]$$

We can get the result directly from the One Sample t-test package. And the 99% Confidence Interval for the average potency is [ 24.35735, 25.83599].

One Sample t-test

```
data:  potencies
t = 0.3604, df = 29, p-value = 0.7212
alternative hypothesis: true mean is not equal to 25
99 percent confidence interval:
 24.35735 25.83599
sample estimates:
mean of x
 25.09667
```

- (d) **Based on your results for part (c), can we conclude that the average potency is 25 mg as advertised?** Based on the results from part (c), we have that the  $p\text{-value} = 0.7212 > 0.01$ , hence we do not have enough information to reject  $H_0$ . Therefore we may conclude that the average potency is 25 mg as advertised. Moreover,  $25 \in [24.35735, 25.83599]$  which confirms our conclusion.

□

## Problem 2

WLabor dataset

*Solution.* (a) **Compute the difference scores between percentages for each year and create a stem-and-leaf plot for these difference scores.**

- The difference scores  $\text{difference} = \text{Year}_{68} - \text{Year}_{72}$ :

```
> data
      City Year_68 Year_72 difference
1      N.Y.   0.42   0.45    -0.03
2       L.A.   0.50   0.50     0.00
3    Chicago   0.52   0.52     0.00
4 Philadelphia   0.45   0.45     0.00
5    Detroit   0.43   0.46    -0.03
6 San Francisco   0.55   0.55     0.00
7      Boston   0.45   0.60    -0.15
8       Pitt.   0.34   0.49    -0.15
9   St. Louis   0.45   0.35     0.10
10 Connecticut   0.54   0.55    -0.01
11 Wash., D.C.   0.42   0.52    -0.10
12      Cinn.   0.51   0.53    -0.02
13  Baltimore   0.49   0.57    -0.08
```

14	Newark	0.54	0.53	0.01
15	Minn/St. Paul	0.50	0.59	-0.09
16	Buffalo	0.58	0.64	-0.06
17	Houston	0.49	0.50	-0.01
18	Patterson	0.56	0.57	-0.01
19	Dallas	0.63	0.64	-0.01

- **Stem-and-leaf plot for these difference scores:**

The decimal point is 1 digit(s) to the left of the |

```

-1 | 550
-0 | 9863321111
 0 | 00001
 1 | 0

```

- (b) **Assess the normality of these difference scores.** The following Shapiro-Wilk normality indicates that the p-value is  $0.04503 < 0.05$ , then reject the  $H_0$ . Therefore, we may conclude that this data does not obey normal distribution. And the corresponding Normal QQ plot can be found in Figure.2 which confirms our conclusion.

Shapiro-Wilk normality test

```

data:  difference
W = 0.89814, p-value = 0.04503

```



Figure 2: Normal QQ plot.

- (c) Based upon these results, use either the Wilcoxon Signed Ranks test or the t-test to determine whether there is a difference between the average percentages in 1968 and the average percentages in 1972. Use  $\alpha = 0.05$ .

- **Wilcoxon Signed Ranks test:** The Wilcoxon signed rank test indicates that the p-value is  $0.01324 < 0.05$ , hence reject  $H_0$ . Therefore we may conclude that alternative hypothesis is valid, i.e. **there is a difference** between the average percentages in 1968 and the average percentages in 1972.

Wilcoxon signed rank test with continuity correction

```
data:  diffence
V = 16, p-value = 0.01324
alternative hypothesis: true location is not equal to 0
95 percent confidence interval:
 -0.08004133 -0.01002685
sample estimates:
(pseudo)median
 -0.04498224
```

- **t-test:** Similarly, the Paired t-test indicates that the p-value is  $0.02435 < 0.05$ , hence reject  $H_0$ . Therefore we may conclude that alternative hypothesis is valid, i.e. **there is a difference** between the average percentages in 1968 and the average percentages in 1972.

Paired t-test

```
data:  Year_68 and Year_72
t = -2.4577, df = 18, p-value = 0.02435
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
 -0.062478527 -0.004889895
sample estimates:
mean of the differences
 -0.03368421
```

- (d) **If you decide that the mean percentages differ, estimate this difference with a 95% Confidence Interval.** Based on the results from part (c), the 99% Confidence Interval of the difference is  **$[-0.08004133, -0.01002685]$**  for Wilcoxon signed rank test, and the 99% Confidence Interval of the difference is  **$[-0.062478527, -0.004889895]$**  for t-test .

□

### Problem 3

Weight dataset

*Solution.* (a) **Create a stacked histogram of the data from each of the two therapies, as well as side-by-side box-and-whisker plots.** The histogram and boxplot of the data can be found in Figure.3.

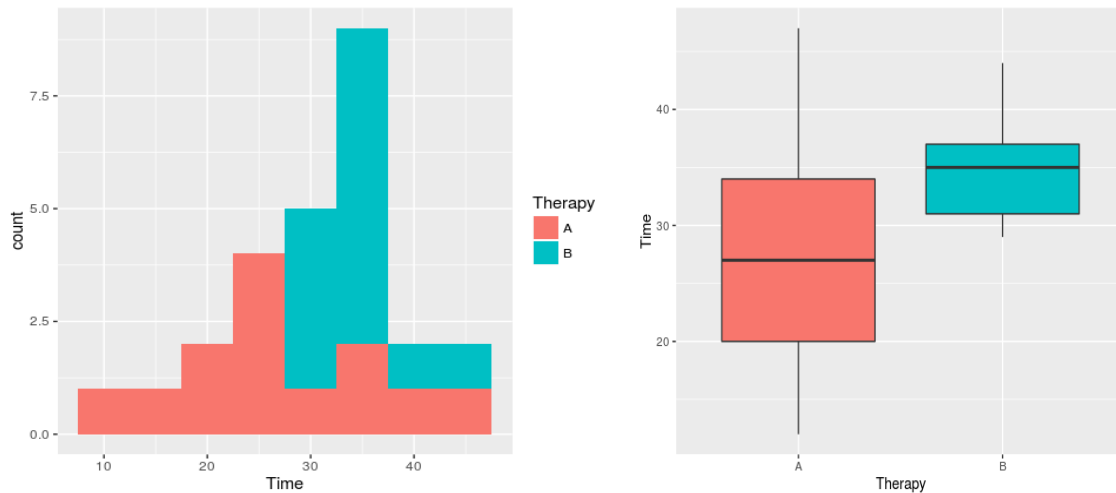


Figure 3: Histogram and boxplot of the data from each of the two therapies.

(b) Assess the normality of the data from each of the two therapies.

- **For therapy A:** The following Shapiro-Wilk normality indicates that the p-value is  $0.7462 > 0.05$ . Hence we do not have enough information to reject the  $H_0$ . Therefore, we may conclude that this data obeys normal distribution. And the corresponding Normal QQ plot can be found in Figure.4 which confirms our conclusion.

Shapiro-Wilk normality test

```
data: group_a
W = 0.95952, p-value = 0.7462
```

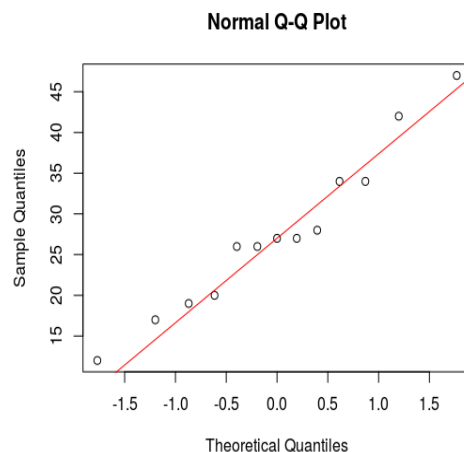


Figure 4: Normal QQ plot.

- **For therapy B:** The following Shapiro-Wilk normality indicates that the p-value is  $0.4024 > 0.05$ . Hence we do not have enough information to reject the  $H_0$ . Therefore, we may conclude that

this data obeys normal distribution. And the corresponding Normal QQ plot can be found in Figure.4 which confirms our conclusion.

Shapiro-Wilk normality test

```
data: group_b
W = 0.93559, p-value = 0.4024
```

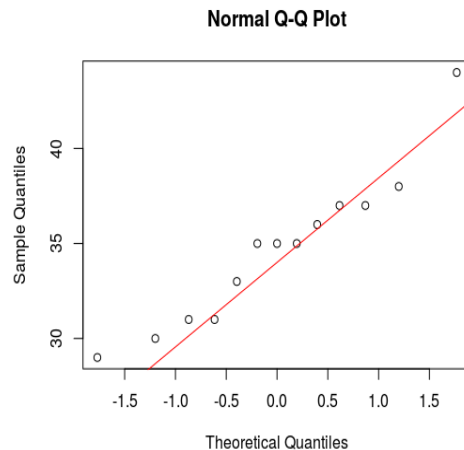


Figure 5: Normal QQ plot.

(c) **Provide a 95% Confidence Interval for the mean times for each of the two therapies.**

- **For therapy A:** From the following One Sample t-test, we get the 95% Confidence Interval for the mean times for each of the therapy A is [21.67638, 33.55439]

One Sample t-test

```
data: group_a
t = 10.131, df = 12, p-value = 3.111e-07
alternative hypothesis: true mean is not equal to 0
95 percent confidence interval:
 21.67638 33.55439
sample estimates:
mean of x
 27.61538
```

- **For therapy B:** From the following One Sample t-test, we get the 95% Confidence Interval for the mean times for each of the therapy B is [32.25776, 37.12685]

One Sample t-test

```
data: group_b
t = 31.048, df = 12, p-value = 7.835e-13
alternative hypothesis: true mean is not equal to 0
```



95 percent confidence interval:

32.25776 37.12685

sample estimates:

mean of x

34.69231

- (d) **Determine if the variances of the times from the two therapies are equal.** According to the following F test, the **p-value = 0.00426 < 0.05**. Hence reject  $H_0$ . Therefore, we may conclude that the variances of the times from the two therapies **are not equal to each other**.

F test to compare two variances

data: group\_a and group\_b

F = 5.951, num df = 12, denom df = 12, p-value = 0.00426

alternative hypothesis: true ratio of variances is not equal to 1

95 percent confidence interval:

1.815845 19.503164

sample estimates:

ratio of variances

5.951027

- (e) **Test whether the means of the times of the two therapies are equal, using either the t-test or the Wilcoxon Rank Sum test, based on the information you obtained above.**

- **Wilcoxon Signed Ranks test:** The Wilcoxon signed rank test indicates that the p-value is  $0.01185 < 0.05$ , hence reject  $H_0$ . Therefore we may conclude that alternative hypothesis is valid, i.e. **the means of the times of the two therapies are not equal to each other**.

Wilcoxon rank sum test with continuity correction

data: group\_a and group\_b

W = 35, p-value = 0.01185

alternative hypothesis: true location shift is not equal to 0

95 percent confidence interval:

-13.999970 -2.000003

sample estimates:

difference in location

-8.000047

- **t-test:** Similarly, the Welch Two Sample t-test indicates that the p-value is  $0.02885 < 0.05$ , hence reject  $H_0$ . Therefore we may conclude that alternative hypothesis is valid, i.e. the means of the times of the two therapies **are not equal to each other**.

Welch Two Sample t-test

data: group\_a and group\_b

t = -2.4023, df = 15.922, p-value = 0.02885

alternative hypothesis: true difference in means is not equal to 0

95 percent confidence interval:

```

-13.3244968 -0.8293494
sample estimates:
mean of x mean of y
27.61538 34.69231

```

- (f) **If you decide that the means of these two therapies are not equal, estimate this difference with a 95% Confidence Interval.** Based on the results from part (d), the 99% Confidence Interval of the difference is `[-13.999970 , -2.000003]` for Wilcoxon signed rank test, and the 99% Confidence Interval of the difference is `[-13.3244968 , -0.8293494]` for t-test .

□

## Problem 4

Weight dataset

- Solution.* (a) **Determine whether clearance to return to work is independent of employee type.** Since the values of each cell are larger than 5, hence we can use  $\chi^2$  test to test it. The following Pearson's Chi-squared test indicates that the p-value =  $6.997e-06 < 0.01$ , hence reject  $H_0$ : Variable A and Variable B are independent. Therefore, we may claim that the clearance to return to work is **dependent on** the employee type.

Pearson's Chi-squared test with Yates' continuity correction

```

data:  ctbl
X-squared = 20.194, df = 1, p-value = 6.997e-06

```

- (b) **Estimate the proportion of salaried workers granted clearance.** The estimated proportion of salaried workers granted clearance is `0.597561` with the 99% Confidence Interval: `[0.4499513 , 0.7299512]`.

1-sample proportions test with continuity correction

```

data:  49 out of 49 + 33, null probability 0.5
X-squared = 2.7439, df = 1, p-value = 0.09763
alternative hypothesis: true p is not equal to 0.5
99 percent confidence interval:
 0.4499513 0.7299512
sample estimates:
      p
0.597561

```

- (c) **Estimate the proportion of wage-earning workers granted clearance.** The estimated proportion of wage-earning workers granted clearance is `0.8731343` with the 99% Confidence Interval: `[0.7767396 , 0.9326389]`.

1-sample proportions test with continuity correction

```
data: 117 out of 117 + 17, null probability 0.5
X-squared = 73.142, df = 1, p-value < 2.2e-16
alternative hypothesis: true p is not equal to 0.5
99 percent confidence interval:
 0.7767396 0.9326389
sample estimates:
      p
0.8731343
```

- (d) **Estimate the difference in these two proportions.** The estimated difference is  $0.5975610 - 0.8731343 = -0.2755733$  with the 99% Confidence Interval:  $[-0.4433355, -0.1078112]$

2-sample test for equality of proportions with continuity correction

```
data: ctbl
X-squared = 20.194, df = 1, p-value = 6.997e-06
alternative hypothesis: two.sided
99 percent confidence interval:
 -0.4433355 -0.1078112
sample estimates:
   prop 1    prop 2 
0.5975610 0.8731343
```

□

## Appendix

### R code for Midterm

Listing 1: Source code for problem 1

```
rm(list = ls())
# set the path or environment
#setwd("/Users/wenqiangfeng/Dropbox/UTK_Course/Stat537/Midterm/data")
setwd("/home/feng/Dropbox/UTK_Course/Stat537/Midterm/data")
5 ##### problem 1 #####
rawdata = read.table("Potencies.txt") # read text file
potencies = unlist(rawdata)

# (a)
10 stem(potencies)

# (b)
## Perform the test
shapiro.test(potencies);
15 ## Plot using a qqplot
```

```

qqnorm(potencias); qqline(potencias, col = 2)

# (c)
20 t.test(potencias, alternative = c("two.sided"),
      mu = 25, conf.level = 0.99)

##### problem 2 #####
#install.packages("readxl") # CRAN version
25 library(readxl)
#install.packages("moments")
rawdata = read_excel("WLabor.xlsx", sheet = 1)
attach(rawdata)

30 # (a)
diffence = Year_68-Year_72
data = cbind(rawdata,diffence)
data
stem(diffence)

35 # (b)
## Perform the test
shapiro.test(diffence);

40 ## Plot using a qqplot
qqnorm(diffence); qqline(diffence, col = 2)

# (c)
wilcox.test(diffence, conf.int = T,
45 alternative="two.sided", conf.level = 0.95)

t.test(Year_68, Year_72, alternative="two.sided",
      paired = TRUE, conf.level = 0.95)

50 ##### problem 3 #####
weight = read.table("Weight.dat", header = TRUE) # read text file
attach(weight)
library(ggplot2)

55 # (a)
# Overlaid histograms with means
ggplot(weight, aes(x=Time, fill=Therapy)) +
  geom_histogram(binwidth = 5)

60 # A basic box plot
#ggplot(weight, aes(x=Therapy, y=Time)) + geom_boxplot()
# The above adds a redundant legend. With the legend removed:
ggplot(weight, aes(x=Therapy, y=Time, fill=Therapy)) + geom_boxplot() +
  guides(fill=FALSE)

65 group_a = weight[c(Therapy=='A'),2]
group_b = weight[c(Therapy=='B'),2]

# (b)

```

```
70 ## Perform the test
shapiro.test(group_a);

## Plot using a qqplot
qqnorm(group_a); qqline(group_a, col = 2)

75 ## Perform the test
shapiro.test(group_b);

## Plot using a qqplot
80 qqnorm(group_b); qqline(group_b, col = 2)

# (c)
t.test(group_a, alternative = c("two.sided"), conf.level = 0.95)
t.test(group_b, alternative = c("two.sided"), conf.level = 0.95)

85 # (d)
var.test(group_a, group_b,
         alternative = c("two.sided"), conf.level = 0.95)

90 # (e)
wilcox.test(group_a, group_b, conf.int = T)

t.test(group_a, group_b, alternative="two.sided", conf.level = 0.95)

95 ##### problem 4 #####
# extra
group = c("Salaried", "Wearning")
granted = c(49, 117)
100 Ngranted = c(33, 17)
data = data.frame(granted, Ngranted)
data

ctbl=cbind(data$granted, data$Ngranted)
105 ctbl

# (a)
chisq.test(ctbl)

110 prop.table(ctbl)

# (b)
prop.test(49, 49+33)

# (c)
115 prop.test(117, 117+17)

# (d)
prop.test(ctbl, alternative="two.sided", conf.level = 0.99)
```